

$$8.16 \quad S^2 = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1} \quad (\text{the sample variance})$$

$$\sum_{i=1}^n (x_i - \bar{x})^2 = \sum_{i=1}^n (x_i^2 - 2x_i\bar{x} + \bar{x}^2) = \sum_{i=1}^n x_i^2 - 2\bar{x} \cdot \underbrace{\sum_{i=1}^n x_i}_{\textcircled{1}} + \underbrace{\sum_{i=1}^n \bar{x}^2}_{\textcircled{2}}$$

Term ① is  $= 2\bar{x} \cdot n\bar{x}$  since  $n\bar{x} = \sum_{i=1}^n x_i$

Term ② is  $= n\bar{x}^2$  so  $\sum (x_i - \bar{x})^2 = \sum_{i=1}^n x_i^2 - 2n\bar{x}^2 + n\bar{x}^2$   
 (since we add  $n \bar{x}^2$ 's)

$$= \sum_{i=1}^n x_i^2 - n\bar{x}^2$$

Thus 
$$S^2 = \frac{\sum_{i=1}^n x_i^2 - n\bar{x}^2}{n-1} = \frac{\sum x_i^2}{n-1} - \frac{n\bar{x}^2}{n-1}$$

8.18 Use mgf's.  $X_i \sim \chi^2(\nu_i) = \Gamma\left(\frac{\nu_i}{2}, 2\right)$

$$M_{X_i}(t) = (1-2t)^{-\nu_i/2} \text{ for } i=1, 2, \dots, n.$$

$$\begin{aligned} \text{So } M_Y(t) &= M_{X_1}(t) \cdot M_{X_2}(t) \cdots M_{X_n}(t) \\ &= (1-2t)^{-\nu_1/2} \cdots (1-2t)^{-\nu_n/2} \\ &= (1-2t)^{-(\nu_1 + \nu_2 + \cdots + \nu_n)/2} \end{aligned}$$

which is the mgf of a  $\chi^2(\nu_1 + \nu_2 + \cdots + \nu_n)$   
r.v.

8.19 If  $X_1 \sim \chi^2(\nu_1)$  and  $X_2, X_1$  are independent  
with  $X_1 + X_2 \sim \chi^2(\nu)$ , then

$$M_{X_1}(t) \cdot M_{X_2}(t) = M_{X_1 + X_2}(t) \Rightarrow$$

$$(1-2t)^{-\nu_1/2} M_{X_2}(t) = (1-2t)^{-\nu/2}$$

$$\text{So } M_{X_2}(t) = \frac{(1-2t)^{-\nu/2}}{(1-2t)^{-\nu_1/2}} = (1-2t)^{-(\nu - \nu_1)/2}$$

$$\Rightarrow X_2 \sim \chi^2(\nu - \nu_1)$$

8.20.

Observe that  $X_i - \mu = (X_i - \bar{X}) + (\bar{X} - \mu)$

$$\begin{aligned} \text{So } \sum_{i=1}^n (X_i - \mu)^2 &= \sum \left[ (X_i - \bar{X})^2 + 2(X_i - \bar{X})(\bar{X} - \mu) + (\bar{X} - \mu)^2 \right] \\ &= \sum_1^n (X_i - \bar{X})^2 + 2(\bar{X} - \mu) \sum_1^n (X_i - \bar{X}) + \sum_1^n (\bar{X} - \mu)^2 \end{aligned}$$

Note that the last term is just the sum of  $n$   $(\bar{X} - \mu)^2$ 's so equals  $n(\bar{X} - \mu)^2$ . The middle

term is 0 since  $\sum_{i=1}^n (X_i - \bar{X}) = \sum_{i=1}^n X_i - \sum_{i=1}^n \bar{X}$   
 $= \sum_{i=1}^n X_i - n\bar{X} = \sum_{i=1}^n X_i - \sum_{i=1}^n \bar{X}$

Thus  $\sum (X_i - \mu)^2 = \sum (X_i - \bar{X})^2 + n(\bar{X} - \mu)^2$ .

(There are many other ways to do this!)

8.46 If  $X_1, \dots, X_n \sim \text{Unif}(0,1]$  the common d.f. is  $F(x) = \begin{cases} x, & 0 < x < 1 \\ 0, & x < 0 \\ 1, & x > 1 \end{cases}$

and common density is  $f(x) = \begin{cases} 1, & 0 < x < 1 \\ 0, & \text{else} \end{cases}$

So the density of  $Y_1$  is  $n(1-F(y))^{n-1}f(y) = \begin{cases} n(1-y)^{n-1}, & 0 < y < 1 \\ 0, & \text{else} \end{cases}$

$$\text{Thus } EY_1 = \int_0^1 y \cdot n(1-y)^{n-1} dy = \frac{1}{n+1}$$

$$EY_1^2 = \int_0^1 y^2 n(1-y)^{n-1} dy = \frac{2}{(n+1)(n+2)}$$

$$\text{Var}(Y_1) = E(Y_1^2) - (EY_1)^2 = \frac{n}{(n+1)^2(n+2)}$$

The easy way to do these integrals is to substitute  $u = 1 - y$ . Then  $du = -dy$  and  $y = 1 - u$  so

$$\begin{aligned} \text{e.g. } \int_0^1 y^2 (1-y)^{n-1} dy &= - \int_1^0 (1-u)^2 u^{n-1} du \\ &= \int_0^1 (1-2u+u^2) u^{n-1} du \\ &= \int_0^1 u^{n-1} du - 2 \int_0^1 u^n du + \int_0^1 u^{n+1} du \\ &= \frac{1}{n} - \frac{2}{n+1} + \frac{1}{n+2} \end{aligned}$$

$$\text{so } EY_1^2 = n \left( \frac{1}{n} - \frac{2}{n+1} + \frac{1}{n+2} \right) = \frac{2}{(n+1)(n+2)}$$

8.67 If  $X_1, \dots, X_{200} \sim \text{Unif}[24, 48]$

then  $\mu = EX_1 = \frac{48+24}{2} = 36$  and  $\sigma^2 = \text{Var}(X_1) = \frac{24^2}{12} = 48$

so  $\sigma = \sqrt{48}$  and

$$P(\bar{X} < 35) = P\left( \frac{\bar{X} - \mu}{\sigma/\sqrt{200}} < \frac{35 - 36}{\sqrt{48}/\sqrt{200}} \right)$$

$$\text{by CLT } \approx P\left( Z < \frac{-1}{\sqrt{\frac{48}{200}}} \right) \approx P(Z < -2.041)$$

$N(0, 1)$

$$= .0206$$